

NAG Library Routine Document

G02MBF

Note: before using this routine, please read the Users' Note for your implementation to check the interpretation of *bold italicised* terms and other implementation-dependent details.

1 Purpose

G02MBF performs Least Angle Regression (LARS), forward stagewise linear regression or Least Absolute Shrinkage and Selection Operator (LASSO) using cross-product matrices.

2 Specification

```

SUBROUTINE G02MBF (MTYPE, PRED, INTCPT, N, M, DTD, LDDTD, ISX, LISX,      &
                  DTY, YTY, MNSTEP, IP, NSTEP, B, LDB, FITSUM, ROPT,    &
                  LROPT, IFAIL)
INTEGER           MTYPE, PRED, INTCPT, N, M, LDDTD, ISX(LISX), LISX,    &
                  MNSTEP, IP, NSTEP, LDB, LROPT, IFAIL
REAL (KIND=nag_wp) DTD(LDDTD,*), DTY(M), YTY, B(LDB,*),              &
                  FITSUM(6,MNSTEP+1), ROPT(LROPT)

```

3 Description

G02MBF implements the LARS algorithm of Efron *et al.* (2004) as well as the modifications needed to perform forward stagewise linear regression and fit LASSO and positive LASSO models.

Given a vector of n observed values, $y = \{y_i : i = 1, 2, \dots, n\}$ and an $n \times p$ design matrix X , where the j th column of X , denoted x_j , is a vector of length n representing the j th independent variable x_j , standardized such that $\sum_{i=1}^n x_{ij} = 0$, and $\sum_{i=1}^n x_{ij}^2 = 1$ and a set of model parameters β to be estimated from the observed values, the LARS algorithm can be summarised as:

1. Set $k = 1$ and all coefficients to zero, that is $\beta = 0$.
2. Find the variable most correlated with y , say x_{j_1} . Add x_{j_1} to the 'most correlated' set \mathcal{A} . If $p = 1$ go to 8.
3. Take the largest possible step in the direction of x_{j_1} (i.e., increase the magnitude of β_{j_1}) until some other variable, say x_{j_2} , has the same correlation with the current residual, $y - x_{j_1}\beta_{j_1}$.
4. Increment k and add x_{j_k} to \mathcal{A} .
5. If $|\mathcal{A}| = p$ go to 8.
6. Proceed in the 'least angle direction', that is, the direction which is equiangular between all variables in \mathcal{A} , altering the magnitude of the parameter estimates of those variables in \mathcal{A} , until the k th variable, x_{j_k} , has the same correlation with the current residual.
7. Go to 4.
8. Let $K = k$.

As well as being a model selection process in its own right, with a small number of modifications the LARS algorithm can be used to fit the LASSO model of Tibshirani (1996), a positive LASSO model, where the independent variables enter the model in their defined direction, forward stagewise linear regression (Hastie *et al.* (2001)) and forward selection (Weisberg (1985)). Details of the required modifications in each of these cases are given in Efron *et al.* (2004).

The LASSO model of Tibshirani (1996) is given by

$$\underset{\alpha, \beta_k \in \mathbb{R}^p}{\text{minimize}} \|y - \alpha - X^T \beta_k\|^2 \quad \text{subject to} \quad \|\beta_k\|_1 \leq t_k$$

for all values of t_k , where $\alpha = \bar{y} = n^{-1} \sum_{i=1}^n y_i$. The positive LASSO model is the same as the standard LASSO model, given above, with the added constraint that

$$\beta_{kj} \geq 0, \quad j = 1, 2, \dots, p.$$

Unlike the standard LARS algorithm, when fitting either of the LASSO models, variables can be dropped as well as added to the set \mathcal{A} . Therefore the total number of steps K is no longer bounded by p .

Forward stagewise linear regression is an iterative procedure of the form:

1. Initialize $k = 1$ and the vector of residuals $r_0 = y - \alpha$.
2. For each $j = 1, 2, \dots, p$ calculate $c_j = x_j^T r_{k-1}$. The value c_j is therefore proportional to the correlation between the j th independent variable and the vector of previous residual values, r_k .
3. Calculate $j_k = \underset{j}{\operatorname{argmax}} |c_j|$, the value of j with the largest absolute value of c_j .
4. If $|c_{j_k}| < \epsilon$ then go to 7.
5. Update the residual values, with

$$r_k = r_{k-1} + \delta \operatorname{sign}(c_{j_k}) x_{j_k}$$

where δ is a small constant and $\operatorname{sign}(c_{j_k}) = -1$ when $c_{j_k} < 0$ and 1 otherwise.

6. Increment k and go to 2.
7. Set $K = k$.

If the largest possible step were to be taken, that is $\delta = |c_{j_k}|$ then forward stagewise linear regression reverts to the standard forward selection method as implemented in G02EEF.

The LARS procedure results in K models, one for each step of the fitting process. In order to aid in choosing which is the most suitable Efron *et al.* (2004) introduced a C_p -type statistic given by

$$C_p^{(k)} = \frac{\|y - X^T \beta_k\|^2}{\sigma^2} - n + 2\nu_k,$$

where ν_k is the approximate degrees of freedom for the k th step and

$$\sigma^2 = \frac{n - y^T y}{\nu_K}.$$

One way of choosing a model is therefore to take the one with the smallest value of $C_p^{(k)}$.

4 References

- Efron B, Hastie T, Johnstone I and Tibshirani R (2004) Least Angle Regression *The Annals of Statistics* (Volume 32) **2** 407–499
- Hastie T, Tibshirani R and Friedman J (2001) *The Elements of Statistical Learning: Data Mining, Inference and Prediction* Springer (New York)
- Tibshirani R (1996) Regression Shrinkage and Selection via the Lasso *Journal of the Royal Statistical Society, Series B (Methodological)* (Volume 58) **1** 267–288
- Weisberg S (1985) *Applied Linear Regression* Wiley

5 Parameters

- 1: MTYPE – INTEGER *Input*
On entry: indicates the type of model to fit.
 MTYPE = 1
 LARS is performed.
 MTYPE = 2
 Forward linear stagewise regression is performed.
 MTYPE = 3
 LASSO model is fit.
 MTYPE = 4
 A positive LASSO model is fit.
Constraint: MTYPE = 1, 2, 3 or 4.
- 2: PRED – INTEGER *Input*
On entry: indicates the type of preprocessing to perform on the cross-products involving the independent variables, i.e., those supplied in DTD and DTY.
 PRED = 0
 No preprocessing is performed.
 PRED = 2
 Each independent variable is normalized, with the j th variable scaled by $1/\sqrt{x_j^T x_j}$. The scaling factor used by variable j is returned in B(j , NSTEP + 1).
Constraint: PRED = 0 or 2.
- 3: INTCPT – INTEGER *Input*
On entry: indicates the type of data preprocessing that was perform on the dependent variable, y , prior to calling this routine.
 INTCPT = 0
 No preprocessing was performed.
 INTCPT = 1
 The dependent variable, y , was mean centered.
Constraint: INTCPT = 0 or 1.
- 4: N – INTEGER *Input*
On entry: n , the number of observations.
Constraint: $N \geq 1$.
- 5: M – INTEGER *Input*
On entry: m , the total number of independent variables.
Constraint: $M \geq 1$.
- 6: DTD(LDDTD, *) – REAL (KIND=nag_wp) array *Input*
Note: the second dimension of the array DTD must be at least $M(M + 1)/2$ if LDDTD = 1, and at least M otherwise.
On entry: $D^T D$, the cross-product matrix, which along with ISX, defines the design matrix cross-product $X^T X$.

If $LDDTD = 1$, $DTD(1, i \times (i - 1)/2 + j)$ must contain the cross-product of the i th and j th variable, for $i = 1, 2, \dots, M$ and $j = 1, 2, \dots, M$. That is the cross-product stacked by columns as returned by G02BUF, for example.

Otherwise $DTD(i, j)$ must contain the cross-product of the i th and j th variable, for $i = 1, 2, \dots, M$ and $j = 1, 2, \dots, M$. It should be noted that, even though $D^T D$ is symmetric, the full matrix must be supplied.

The matrix specified in DTD must be a valid cross-products matrix.

7: LDDTD – INTEGER *Input*

On entry: the first dimension of the array DTD as declared in the (sub)program from which G02MBF is called.

Constraint: $LDDTD = 1$ or $LDDTD \geq M$.

8: ISX(LISX) – INTEGER array *Input*

On entry: indicates which independent variables from DTD will be included in the design matrix, X .

If $LISX = 0$, all variables are included in the design matrix and ISX is not referenced.

If $LISX = M$, for $j = 1, 2, \dots, M$ when $ISX(j)$ must be set as follows:

$$ISX(j) = 1$$

To indicate that the j th variable, as supplied in DTD, is included in the design matrix;

$$ISX(j) = 0$$

To indicate that the j th variable, as supplied in DTD, is not included in the design matrix;

$$\text{and } p = \sum_{j=1}^m ISX(j).$$

Constraint: if $LISX = M$, $ISX(j) = 0$ or 1 and at least one value of $ISX(j) \neq 0$, for $j = 1, 2, \dots, M$.

9: LISX – INTEGER *Input*

On entry: length of the ISX array.

Constraint: $LISX = 0$ or M .

10: DTY(M) – REAL (KIND=nag_wp) array *Input*

On entry: $D^T y$, the cross-product between the dependent variable, y , and the independent variables D .

11: YTY – REAL (KIND=nag_wp) *Input*

On entry: $y^T y$, the sums of squares of the dependent variable.

Constraint: $YTY > 0.0$.

12: MNSTEP – INTEGER *Input*

On entry: the maximum number of steps to carry out in the model fitting process.

If $MTYPE = 1$, i.e., a LARS is being performed, the maximum number of steps the algorithm will take is $\min(p, n)$ if $INTCPT = 0$, otherwise $\min(p, n - 1)$.

If $MTYPE = 2$, i.e., a forward linear stagewise regression is being performed, the maximum number of steps the algorithm will take is likely to be several orders of magnitude more and is no longer bound by p or n .

If $MTYPE = 3$ or 4 , i.e., a LASSO or positive LASSO model is being fit, the maximum number of steps the algorithm will take lies somewhere between that of the LARS and forward linear stagewise regression, again it is no longer bound by p or n .

Constraint: $MNSTEP \geq 1$.

- 13: IP – INTEGER *Output*
On exit: p , number of parameter estimates.
 If $LISX = 0$, $p = M$, i.e., the number of variables in DTD.
 Otherwise p is the number of nonzero values in ISX.
- 14: NSTEP – INTEGER *Output*
On exit: K , the actual number of steps carried out in the model fitting process.
- 15: B(LDB, *) – REAL (KIND=nag_wp) array *Output*
Note: the second dimension of the array B must be at least $MNSTEP + 1$.
On exit: β the parameter estimates, with $B(j, k) = \beta_{kj}$, the parameter estimate for the j th variable, $j = 1, 2, \dots, p$ at the k th step of the model fitting process, $k = 1, 2, \dots, NSTEP$.
 By default, when $PRED = 2$ the parameter estimates are rescaled prior to being returned. If the parameter estimates are required on the normalized scale, then this can be overridden via ROPT.
 The values held in the remaining part of B depend on the type of preprocessing performed.
 If $PRED = 0$ $B(j, NSTEP + 1) = 1$,
 if $PRED = 2$ $B(j, NSTEP + 1) = 1/\sqrt{x_j^T x_j}$,
 for $j = 1, 2, \dots, p$.
- 16: LDB – INTEGER *Input*
On entry: the first dimension of the array B as declared in the (sub)program from which G02MBF is called.
Constraint: $LDB \geq p$, where p is the number of parameter estimates as described in IP.
- 17: FITSUM(6, $MNSTEP + 1$) – REAL (KIND=nag_wp) array *Output*
On exit: summaries of the model fitting process. When $k = 1, 2, \dots, NSTEP$
- FITSUM(1, k)
 $\|\beta_k\|_1$, the sum of the absolute values of the parameter estimates for the k th step of the modelling fitting process. If $PRED = 2$, the scaled parameter estimates are used in the summation.
- FITSUM(2, k)
 RSS_k , the residual sums of squares for the k th step, where $RSS_k = \|y - X^T \beta_k\|^2$.
- FITSUM(3, k)
 ν_k , approximate degrees of freedom for the k th step.
- FITSUM(4, k)
 $C_p^{(k)}$, a C_p -type statistic for the k th step, where $C_p^{(k)} = \frac{RSS_k}{\sigma^2} - n + 2\nu_k$.
- FITSUM(5, k)
 \hat{C}_k , correlation between the residual at step $k - 1$ and the most correlated variable not yet in the active set \mathcal{A} , where the residual at step 0 is y .
- FITSUM(6, k)
 $\hat{\gamma}_k$, the step size used at step k .

In addition

FITSUM(1, NSTEP + 1)
0.

FITSUM(2, NSTEP + 1)

RSS_0 , the residual sums of squares for the null model, where $RSS_0 = y^T y$.

FITSUM(3, NSTEP + 1)

ν_0 , the degrees of freedom for the null model, where $\nu_0 = 0$ if INTCPT = 0 and $\nu_0 = 1$ otherwise.

FITSUM(4, NSTEP + 1)

$C_p^{(0)}$, a C_p -type statistic for the null model, where $C_p^{(0)} = \frac{RSS_0}{\sigma^2} - n + 2\nu_0$.

FITSUM(5, NSTEP + 1)

σ^2 , where $\sigma^2 = \frac{n - RSS_K}{\nu_K}$ and $K = \text{NSTEP}$.

Although the C_p statistics described above are returned when IFAIL = 122 they may not be meaningful due to the estimate σ^2 not being based on the saturated model.

18: ROPT(LROPT) – REAL (KIND=nag_wp) array *Input*

On entry: optional parameters to control various aspects of the LARS algorithm.

The default value will be used for ROPT(i) if LROPT < i , therefore setting LROPT = 0 will use the default values for all optional arguments and ROPT need not be set. The default value will also be used if an invalid value is supplied for a particular argument, for example, setting ROPT(i) = -1 will use the default value for argument i .

ROPT(1)

The minimum step size that will be taken.

Default is $100 \times eps$ is used, where eps is the *machine precision* returned by X02AJF.

ROPT(2)

General tolerance, used amongst other things, for comparing correlations.

Default is ROPT(1).

ROPT(3)

If set to 1 then parameter estimates are rescaled before being returned. If set to 0 then no rescaling is performed. This argument has no effect when PRED = 0.

Default is for the parameter estimates to be rescaled.

Constraints:

ROPT(1) > *machine precision*;

ROPT(2) > *machine precision*.

19: LROPT – INTEGER *Input*

On entry: length of the options array ROPT.

Constraint: $0 \leq \text{LROPT} \leq 3$.

20: IFAIL – INTEGER *Input/Output*

On entry: IFAIL must be set to 0, -1 or 1. If you are unfamiliar with this parameter you should refer to Section 3.3 in the Essential Introduction for details.

For environments where it might be inappropriate to halt program execution when an error is detected, the value -1 or 1 is recommended. If the output of error messages is undesirable, then the value 1 is recommended. Otherwise, because for this routine the values of the output parameters may be useful even if IFAIL \neq 0 on exit, the recommended value is -1. **When the value -1 or 1 is used it is essential to test the value of IFAIL on exit.**

On exit: IFAIL = 0 unless the routine detects an error or a warning has been flagged (see Section 6).

6 Error Indicators and Warnings

If on entry IFAIL = 0 or -1, explanatory error messages are output on the current error message unit (as defined by X04AAF).

Note: G02MBF may return useful information for one or more of the following detected errors or warnings.

Errors or warnings detected by the routine:

IFAIL = 11

On entry, MTYPE = $\langle value \rangle$.
Constraint: MTYPE = 1, 2, 3 or 4.

IFAIL = 21

On entry, PRED = $\langle value \rangle$.
Constraint: PRED = 0 or 2.

IFAIL = 31

On entry, INTCPT = $\langle value \rangle$.
Constraint: INTCPT = 0 or 1.

IFAIL = 41

On entry, N = $\langle value \rangle$.
Constraint: $N \geq 1$.

IFAIL = 51

On entry, M = $\langle value \rangle$.
Constraint: $M \geq 1$.

IFAIL = 61

The cross-product matrix supplied in DTD is not symmetric.

IFAIL = 62

On entry, $DTD(1, \langle value \rangle) = \langle value \rangle$.
Constraint: diagonal elements of $D^T D$ must be positive.
On entry, $i = \langle value \rangle$ and $DTD(i, i) = \langle value \rangle$.
Constraint: diagonal elements of $D^T D$ must be positive.

IFAIL = 71

On entry, LDDTD = $\langle value \rangle$ and M = $\langle value \rangle$.
Constraint: LDDTD = 1 or LDDTD \geq M.

IFAIL = 81

On entry, $ISX(\langle value \rangle) = \langle value \rangle$.
Constraint: $ISX(i) = 0$ or 1 for all i .

IFAIL = 82

On entry, all values of ISX are zero.
Constraint: at least one value of ISX must be nonzero.

IFAIL = 91

On entry, LISX = $\langle value \rangle$ and M = $\langle value \rangle$.
Constraint: LISX = 0 or M.

IFAIL = 111

On entry, YTY = $\langle value \rangle$.
Constraint: YTY > 0.0.

IFAIL = 112

A negative value for the residual sums of squares was obtained. Check the values of DTD, DTY and YTY.

IFAIL = 121

On entry, MNSTEP = $\langle value \rangle$.
Constraint: MNSTEP \geq 1.

IFAIL = 122

Fitting process did not finished in MNSTEP steps. Try increasing the size of MNSTEP and supplying larger output arrays.
All output is returned as documented, up to step MNSTEP, however, σ and the C_p statistics may not be meaningful.

IFAIL = 161

On entry, LDB = $\langle value \rangle$ and M = $\langle value \rangle$.
Constraint: if LISX = 0 then LDB \geq M.

IFAIL = 162

On entry, LDB = $\langle value \rangle$ and p = $\langle value \rangle$.
Constraint: if LISX = M, LDB \geq p .

IFAIL = 171

σ^2 is approximately zero and hence the C_p -type criterion cannot be calculated. All other output is returned as documented.

IFAIL = 172

$\nu_K = n$, therefore sigma has been set to a large value. Output is returned as documented.

IFAIL = 173

Degenerate model, no variables added and NSTEP = 0. Output is returned as documented.

IFAIL = 191

On entry, LROPT = $\langle value \rangle$.
Constraint: $0 \leq$ LROPT \leq 3.

IFAIL = -99

An unexpected error has been triggered by this routine. Please contact NAG.
See Section 3.8 in the Essential Introduction for further information.

IFAIL = -399

Your licence key may have expired or may not have been installed correctly.
See Section 3.7 in the Essential Introduction for further information.

IFAIL = -999

Dynamic memory allocation failed.

See Section 3.6 in the Essential Introduction for further information.

7 Accuracy

Not applicable.

8 Parallelism and Performance

G02MBF is threaded by NAG for parallel execution in multithreaded implementations of the NAG Library.

G02MBF makes calls to BLAS and/or LAPACK routines, which may be threaded within the vendor library used by this implementation. Consult the documentation for the vendor library for further information.

Please consult the X06 Chapter Introduction for information on how to control and interrogate the OpenMP environment used within this routine. Please also consult the Users' Note for your implementation for any additional implementation-specific information.

9 Further Comments

The solution path to the LARS, LASSO and stagewise regression analysis is a continuous, piecewise linear. G02MBF returns the parameter estimates at various points along this path. G02MCF can be used to obtain estimates at different points along the path.

If you have the raw data values, that is D and y , then G02MAF can be used instead of G02MBF.

10 Example

This example performs a LARS on a simulated dataset with 20 observations and 6 independent variables.

The example uses G02BUF to get the cross-products of the augmented matrix $[D \ y]$. The first $m(m+1)/2$ elements of the (column packed) cross-products matrix returned therefore contain the elements of $D^T D$, the next m elements contain $D^T y$ and the last element $y^T y$.

10.1 Program Text

```

Program g02mbfe

!      G02MBF Example Program Text

!      Mark 25 Release. NAG Copyright 2014.

!      .. Use Statements ..
      Use nag_library, Only: g02buf, g02mbf, nag_wp
!      .. Implicit None Statement ..
      Implicit None
!      .. Parameters ..
      Integer, Parameter          :: nin = 5, nout = 6
!      .. Local Scalars ..
      Real (Kind=nag_wp)          :: sw
      Integer                      :: i, ifail, intcpt, ip, k, ldb, ldtd, &
                                   lisx, lropt, m, mnstep, mtype, n,      &
                                   nstep, pm, pm2, pred
      Character (10)              :: mean
!      .. Local Arrays ..
      Real (Kind=nag_wp), Allocatable :: b(:,,:), dtd(:,,:), dy(:,,:),      &
                                   fitsum(:,,:), ropt(:), wmean(:)
      Real (Kind=nag_wp)          :: wt(1)
      Integer, Allocatable        :: isx(:)
!      .. Intrinsic Procedures ..

```

```

      Intrinsic                :: count, floor, max, repeat
!   .. Executable Statements ..

!   .. Executable Statements ..
Write (nout,*) 'G02MBF Example Program Results'
Write (nout,*)

!   Skip heading in data file
Read (nin,*)

!   Read in the problem size
Read (nin,*) n, m

!   Read in the model specification
Read (nin,*) mtype, pred, intcpt, mnstep, lisx

!   Read in variable inclusion flags (if specified) and calculate IP
Allocate (isx(lisx))
If (lisx==m) Then
  Read (nin,*) isx(1:lisx)
  ip = count(isx(1:m)==1)
Else
  ip = m
End If

!   Optional arguments (using defaults)
lropt = 0
Allocate (ropt(lropt))

!   Read in the augmented matrix [D y] and calculate cross-product matrices
!   (NB: Datasets with a large number of observations can be split into
!   blocks with the resulting cross-product matrices being combined
!   using G02BZF)
Allocate (dy(n,m+1))
Read (nin,*)(dy(i,1:m),dy(i,m+1),i=1,n)

pm = m*(m+1)/2
pm2 = (m+1)*(m+1)/2

!   We are calculating the cross-product matrix using G02BUF
!   which returns it in packed storage
lddtd = 1

!   Calculate the cross-product matrices
Allocate (wmean(m+1),dtd(1,pm2))
If (intcpt==1) Then
  mean = 'Mean'
Else
  mean = 'Zero'
End If
ifail = 0
Call g02buf(mean,'UnWeighted',n,m+1,dy,n,wt,sw,wmean,dtd(1,:),ifail)
!   The first PM elements of DTD(1,:) contain the cross-products of D
!   elements DTD(1,PM+1:PM2-1) contains cross-product of D with y and
!   DTD(1,PM2) contains cross-product of y with itself

!   Allocate output arrays
ldb = ip
Allocate (b(ldb,mnstep+1),fitsum(6,mnstep+1))

!   Call the model fitting routine
ifail = -1
Call g02mbf(mtype,pred,intcpt,n,m,dtd,lddtd,isx,lisx,dtd(1,pm+1:pm2-1), &
  dtd(1,pm2),mnstep,ip,nstep,b,ldb,fitsum,ropt,lropt,ifail)
If (ifail/=0) Then
  If (ifail/=112 .And. ifail/=161 .And. ifail/=162 .And. ifail/=163) &
  Then
!     IFAIL = 112, 161, 162 and 163 are warnings, so no need to terminate
!     if they occur
  Go To 100
  End If

```

```

End If

! Display the parameter estimates
Write (nout,*) ' Step ', repeat(' ',max((ip-2),0)*5), &
' Parameter Estimate'
Write (nout,*) repeat('-',5+ip*10)
Do k = 1, nstep
  Write (nout,99998) k, b(1:ip,k)
End Do
Write (nout,*)
Write (nout,99999) 'alpha: ', wmean(m+1)
Write (nout,*)
Write (nout,*) &
' Step      Sum      RSS      df      Cp      Ck      Step Size'
Write (nout,*) repeat('-',64)
Do k = 1, nstep
  Write (nout,99997) k, fitsum(1:2,k), floor(fitsum(3,k)+0.5_nag_wp), &
  fitsum(4:6,k)
End Do
Write (nout,*)
Write (nout,99999) 'sigma^2: ', fitsum(5,nstep+1)

100 Continue
99999 Format (1X,A,F9.3)
99998 Format (2X,I3,10(1X,F9.3))
99997 Format (2X,I3,2(1X,F9.3),1X,I6,1X,3(1X,F9.3))
End Program g02mbfe

```

10.2 Program Data

G02MBF Example Program Data

```

20 6          :: N,M
1 2 1 6 0     :: MTYPE,PRED,INTCPT,MNSTEP,LISX
10.28  1.77  9.69 15.58  8.23 10.44  -46.47
 9.08  8.99 11.53  6.57 15.89 12.58  -35.80
17.98 13.10  1.04 10.45 10.12 16.68  -129.22
14.82 13.79 12.23  7.00  8.14  7.79  -42.44
17.53  9.41  6.24  3.75 13.12 17.08  -73.51
 7.78 10.38  9.83  2.58 10.13  4.25  -26.61
11.95 21.71  8.83 11.00 12.59 10.52  -63.90
14.60 10.09 -2.70  9.89 14.67  6.49  -76.73
 3.63  9.07 12.59 14.09  9.06  8.19  -32.64
 6.35  9.79  9.40 12.79  8.38 16.79  -83.29
 4.66  3.55 16.82 13.83 21.39 13.88  -16.31
 8.32 14.04 17.17  7.93  7.39 -1.09   -5.82
10.86 13.68  5.75 10.44 10.36 10.06  -47.75
 4.76  4.92 17.83  2.90  7.58 11.97   18.38
 5.05 10.41  9.89  9.04  7.90 13.12  -54.71
 5.41  9.32  5.27 15.53  5.06 19.84  -55.62
 9.77  2.37  9.54 20.23  9.33  8.82  -45.28
14.28  4.34 14.23 14.95 18.16 11.03  -22.76
10.17  6.80  3.17  8.57 16.07 15.93  -104.32
 5.39  2.67  6.37 13.56 10.68  7.35  -55.94 :: End of D, Y

```

10.3 Program Results

G02MBF Example Program Results

Step	Parameter Estimate					
1	0.000	0.000	3.125	0.000	0.000	0.000
2	0.000	0.000	3.792	0.000	0.000	-0.713
3	-0.446	0.000	3.998	0.000	0.000	-1.151
4	-0.628	-0.295	4.098	0.000	0.000	-1.466
5	-1.060	-1.056	4.110	-0.864	0.000	-1.948
6	-1.073	-1.132	4.118	-0.935	-0.059	-1.981

alpha: -50.037

Step	Sum	RSS	df	Cp	Ck	Step Size
------	-----	-----	----	----	----	-----------

1	72.446	8929.855	2	13.355	123.227	72.446
2	103.385	6404.701	3	7.054	50.781	24.841
3	126.243	5258.247	4	5.286	30.836	16.225
4	145.277	4657.051	5	5.309	19.319	11.587
5	198.223	3959.401	6	5.016	12.266	24.520
6	203.529	3954.571	7	7.000	0.910	2.198

sigma^2: 304.198

This example plot shows the regression coefficients (β_k) plotted against the scaled absolute sum of the parameter estimates ($\|\beta_k\|_1$).

