NAG Library Routine Document

G02QGF

Note: before using this routine, please read the Users' Note for your implementation to check the interpretation of **bold italicised** terms and other implementation-dependent details.

Note: this routine uses **optional parameters** to define choices in the problem specification and in the details of the algorithm. If you wish to use default settings for all of the optional parameters, you need only read Sections 1 to 10 of this document. If, however, you wish to reset some or all of the settings please refer to Section 11 for a detailed description of the algorithm, to Section 12 for a detailed description of the specification of the optional parameters and to Section 13 for a detailed description of the monitoring information produced by the routine.

1 Purpose

G02QGF performs a multiple linear quantile regression. Parameter estimates and, if required, confidence limits, covariance matrices and residuals are calculated. G02QGF may be used to perform a weighted quantile regression. A simplified interface for G02QGF is provided by G02QFF.

2 Specification

```
SUBROUTINE GO2QGF (SORDER, INTCPT, WEIGHT, N, M, DAT, LDDAT, ISX, IP, Y, & WT, NTAU, TAU, DF, B, BL, BU, CH, RES, IOPTS, OPTS, STATE, INFO, IFAIL)

INTEGER SORDER, N, M, LDDAT, ISX(M), IP, NTAU, IOPTS(*), STATE(*), INFO(NTAU), IFAIL

REAL (KIND=nag_wp) DAT(LDDAT,*), Y(N), WT(*), TAU(NTAU), DF, & B(IP,NTAU), BL(IP,*), BU(IP,*), CH(IP,IP,*), RES(N,*), OPTS(*)

CHARACTER(1) INTCPT, WEIGHT
```

3 Description

Given a vector of n observed values, $y = \{y_i : i = 1, 2, ..., n\}$, an $n \times p$ design matrix X, a column vector, x, of length p holding the ith row of X and a quantile $\tau \in (0, 1)$, G02QGF estimates the p-element vector β as the solution to

$$\underset{\beta \in \mathbb{R}^p}{\text{minimize}} \sum_{i=1}^n \rho_\tau (y_i - x_i^{\mathsf{T}} \beta) \tag{1}$$

where ρ_{τ} is the piecewise linear loss function $\rho_{\tau}(z) = z(\tau - I(z < 0))$, and I(z < 0) is an indicator function taking the value 1 if z < 0 and 0 otherwise. Weights can be incorporated by replacing X and y with WX and Wy respectively, where W is an $n \times n$ diagonal matrix. Observations with zero weights can either be included or excluded from the analysis; this is in contrast to least squares regression where such observations do not contribute to the objective function and are therefore always dropped.

G02QGF uses the interior point algorithm of Portnoy and Koenker (1997), described briefly in Section 11, to obtain the parameter estimates $\hat{\beta}$, for a given value of τ .

Under the assumption of Normally distributed errors, Koenker (2005) shows that the limiting covariance matrix of $\hat{\beta} - \beta$ has the form

$$\Sigma = \frac{\tau(1-\tau)}{n} H_n^{-1} J_n H_n^{-1}$$

where $J_n = n^{-1} \sum_{i=1}^n x_i x_i^{\mathrm{T}}$ and H_n is a function of τ , as described below. Given an estimate of the covariance matrix, $\hat{\Sigma}$, lower $(\hat{\beta}_L)$ and upper $(\hat{\beta}_U)$ limits for an $(100 \times \alpha)\%$ confidence interval can be

calculated for each of the p parameters, via

$$\hat{\beta}_{Li} = \hat{\beta}_i - t_{n-p,(1+\alpha)/2} \sqrt{\hat{\Sigma}_{ii}}, \hat{\beta}_{Ui} = \hat{\beta}_i + t_{n-p,(1+\alpha)/2} \sqrt{\hat{\Sigma}_{ii}}$$

where $t_{n-p,0.975}$ is the 97.5 percentile of the Student's t distribution with n-k degrees of freedom, where k is the rank of the cross-product matrix $X^{T}X$.

Four methods for estimating the covariance matrix, Σ , are available:

(i) Independent, identically distributed (IID) errors

Under an assumption of IID errors the asymptotic relationship for Σ simplifies to

$$\Sigma = \frac{\tau(1-\tau)}{n} (s(\tau))^2 (X^{\mathsf{T}} X)^{-1}$$

where s is the sparsity function. G02QGF estimates $s(\tau)$ from the residuals, $r_i = y_i - x_i^T \hat{\beta}$ and a bandwidth h_n .

(ii) Powell Sandwich

Powell (1991) suggested estimating the matrix H_n by a kernel estimator of the form

$$\hat{H}_n = (nc_n)^{-1} \sum_{i=1}^n K\left(\frac{r_i}{c_n}\right) x_i x_i^{\mathsf{T}}$$

where K is a kernel function and c_n satisfies $\lim_{n\to\infty}c_n\to 0$ and $\lim_{n\to\infty}\sqrt{n}c_n\to \infty$. When the Powell method is chosen, G02QGF uses a Gaussian kernel (i.e., $K=\phi$) and sets

$$c_n = \min(\sigma_r, (q_{r3} - q_{r1})/1.34) \times (\Phi^{-1}(\tau + h_n) - \Phi^{-1}(\tau - h_n))$$

where h_n is a bandwidth, σ_r , q_{r1} and q_{r3} are, respectively, the standard deviation and the 25% and 75% quantiles for the residuals, r_i .

(iii) Hendricks-Koenker Sandwich

Koenker (2005) suggested estimating the matrix H_n using

$$\hat{H}_n = n^{-1} \sum_{i=1}^n \left[\frac{2h_n}{x_i^{\mathsf{T}} \left(\hat{\beta}(\tau + h_n) - \hat{\beta}(\tau - h_n) \right)} \right] x_i x_i^{\mathsf{T}}$$

where h_n is a bandwidth and $\hat{\beta}(\tau + h_n)$ denotes the parameter estimates obtained from a quantile regression using the $(\tau + h_n)$ th quantile. Similarly with $\hat{\beta}(\tau - h_n)$.

(iv) Bootstrap

The last method uses bootstrapping to either estimate a covariance matrix or obtain confidence intervals for the parameter estimates directly. This method therefore does not assume Normally distributed errors. Samples of size n are taken from the paired data $\{y_i, x_i\}$ (i.e., the independent and dependent variables are sampled together). A quantile regression is then fitted to each sample resulting in a series of bootstrap estimates for the model parameters, β . A covariance matrix can then be calculated directly from this series of values. Alternatively, confidence limits, $\hat{\beta}_L$ and $\hat{\beta}_U$, can be obtained directly from the $(1-\alpha)/2$ and $(1+\alpha)/2$ sample quantiles of the bootstrap estimates.

Further details of the algorithms used to calculate the covariance matrices can be found in Section 11. All three asymptotic estimates of the covariance matrix require a bandwidth, h_n . Two alternative methods for determining this are provided:

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(i) Sheather-Hall

$$h_n = \left(\frac{1.5\left(\Phi^{-1}\left(\alpha_b\right)\phi\left(\Phi^{-1}\left(\tau\right)\right)\right)^2}{n\left(2\Phi^{-1}\left(\tau\right)+1\right)}\right)^{\frac{1}{3}}$$

for a user-supplied value α_b ,

(ii) Bofinger

$$h_n = \left(\frac{4.5\left(\phi\left(\Phi^{-1}\left(\tau\right)\right)\right)^4}{n\left(2\Phi^{-1}\left(\tau\right)+1\right)^2}\right)^{\frac{1}{5}}$$

G02QGF allows optional arguments to be supplied via the IOPTS and OPTS arrays (see Section 12 for details of the available options). Prior to calling G02QGF the optional parameter arrays, IOPTS and OPTS must be initialized by calling G02ZKF with OPTSTR set to **Initialize** = G02QGF (see Section 12 for details on the available options). If bootstrap confidence limits are required (**Interval Method** = BOOTSTRAP XY) then one of the random number initialization routines G05KFF (for a repeatable analysis) or G05KGF (for an unrepeatable analysis) must also have been previously called.

4 References

Koenker R (2005) Quantile Regression Econometric Society Monographs, Cambridge University Press, New York

Mehrotra S (1992) On the implementation of a primal-dual interior point method SIAM J. Optim. 2 575-601

Nocedal J and Wright S J (1999) Numerical Optimization Springer Series in Operations Research, Springer, New York

Portnoy S and Koenker R (1997) The Gaussian hare and the Laplacian tortoise: computability of squared-error versus absolute error estimators *Statistical Science* **4** 279–300

Powell J L (1991) Estimation of monotonic regression models under quantile restrictions Nonparametric and Semiparametric Methods in Econometrics Cambridge University Press, Cambridge

5 Arguments

1: SORDER – INTEGER

Input

On entry: determines the storage order of variates supplied in DAT.

Constraint: SORDER = 1 or 2.

2: INTCPT - CHARACTER(1)

Input

On entry: indicates whether an intercept will be included in the model. The intercept is included by adding a column of ones as the first column in the design matrix, X.

INTCPT = 'Y'

An intercept will be included in the model.

INTCPT = 'N'

An intercept will not be included in the model.

Constraint: INTCPT = 'N' or 'Y'.

3: WEIGHT - CHARACTER(1)

Input

On entry: indicates if weights are to be used.

WEIGHT = 'W'

A weighted regression model is fitted to the data using weights supplied in array WT.

WEIGHT = 'U'

An unweighted regression model is fitted to the data and array WT is not referenced.

Constraint: WEIGHT = 'U' or 'W'.

4: N – INTEGER Input

On entry: the total number of observations in the dataset. If no weights are supplied, or no zero weights are supplied or observations with zero weights are included in the model then N=n. Otherwise N=n+ the number of observations with zero weights.

Constraint: $N \ge 2$.

5: M – INTEGER Input

On entry: m, the total number of variates in the dataset.

Constraint: $M \ge 0$.

6: DAT(LDDAT, *) - REAL (KIND=nag wp) array

Input

Note: the second dimension of the array DAT must be at least M if SORDER = 1 and at least N if SORDER = 2.

On entry: the *i*th value for the *j*th variate, for i = 1, 2, ..., N and j = 1, 2, ..., M, must be supplied in

```
DAT(i, j) if SORDER = 1, and DAT(j, i) if SORDER = 2.
```

The design matrix X is constructed from DAT, ISX and INTCPT.

7: LDDAT – INTEGER

Input

On entry: the first dimension of the array DAT as declared in the (sub)program from which G02QGF is called.

Constraints:

```
if SORDER = 1, LDDAT \ge N; otherwise LDDAT \ge M.
```

8: ISX(M) – INTEGER array

Input

On entry: indicates which independent variables are to be included in the model.

```
ISX(j) = 0
```

The jth variate, supplied in DAT, is not included in the regression model.

```
ISX(i) = 1
```

The jth variate, supplied in DAT, is included in the regression model.

Constraints:

```
ISX(j) = 0 or 1, for j = 1, 2, ..., M; if INTCPT = 'Y', exactly IP - 1 values of ISX must be set to 1; if INTCPT = 'N', exactly IP values of ISX must be set to 1.
```

9: IP – INTEGER

Input

On entry: p, the number of independent variables in the model, including the intercept, see INTCPT, if present.

Constraints:

```
1 \le IP < N;
if INTCPT = 'Y', 1 \le IP \le M + 1;
if INTCPT = 'N', 1 \le IP \le M.
```

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10: Y(N) - REAL (KIND=nag wp) array

Input

On entry: y, the observations on the dependent variable.

11: WT(*) - REAL (KIND=nag_wp) array

Input

Note: the dimension of the array WT must be at least N if WEIGHT = 'W'.

On entry: if WEIGHT = 'W', WT must contain the diagonal elements of the weight matrix W. Otherwise WT is not referenced.

When

Drop Zero Weights = YES

If WT(i) = 0.0, the *i*th observation is not included in the model, in which case the effective number of observations, n, is the number of observations with nonzero weights. If **Return Residuals** = YES, the values of RES will be set to zero for observations with zero weights.

Drop Zero Weights = NO

All observations are included in the model and the effective number of observations is N, i.e., n = N.

Constraints:

If WEIGHT = 'W', WT(i) \geq 0.0, for i = 1,2,...,N; The effective number of observations \geq 2.

12: NTAU – INTEGER

Input

On entry: the number of quantiles of interest.

Constraint: NTAU ≥ 1 .

13: TAU(NTAU) – REAL (KIND=nag_wp) array

Input

On entry: the vector of quantiles of interest. A separate model is fitted to each quantile.

Constraint: $\sqrt{\epsilon} < \text{TAU}(j) < 1 - \sqrt{\epsilon}$ where ϵ is the **machine precision** returned by X02AJF, for j = 1, 2, ..., NTAU.

14: DF – REAL (KIND=nag wp)

Output

On exit: the degrees of freedom given by n - k, where n is the effective number of observations and k is the rank of the cross-product matrix $X^{T}X$.

15: B(IP, NTAU) - REAL (KIND=nag_wp) array

Input/Output

On entry: if Calculate Initial Values = NO, B(i, l) must hold an initial estimates for $\hat{\beta}_i$, for i = 1, 2, ..., IP and l = 1, 2, ..., NTAU. If Calculate Initial Values = YES, B need not be set.

On exit: B(i, l), for i = 1, 2, ..., IP, contains the estimates of the parameters of the regression model, $\hat{\beta}$, estimated for $\tau = TAU(l)$.

If INTCPT = 'Y', B(1,l) will contain the estimate corresponding to the intercept and B(i+1,l) will contain the coefficient of the jth variate contained in DAT, where ISX(j) is the ith nonzero value in the array ISX.

If INTCPT = 'N', B(i, l) will contain the coefficient of the *j*th variate contained in DAT, where ISX(j) is the *i*th nonzero value in the array ISX.

16: BL(IP,*) - REAL (KIND=nag wp) array

Output

Note: the second dimension of the array BL must be at least NTAU if Interval Method \neq NONE.

On exit: if Interval Method \neq NONE, BL(i, l) contains the lower limit of an $(100 \times \alpha)\%$ confidence interval for B(i, l), for i = 1, 2, ..., IP and l = 1, 2, ..., NTAU.

If **Interval Method** = NONE, BL is not referenced.

The method used for calculating the interval is controlled by the optional parameters **Interval Method** and **Bootstrap Interval Method**. The size of the interval, α , is controlled by the optional parameter **Significance Level**.

17: BU(IP,*) - REAL (KIND=nag wp) array

Output

Note: the second dimension of the array BU must be at least NTAU if **Interval Method** \neq NONE.

On exit: if Interval Method \neq NONE, BU(i, l) contains the upper limit of an $(100 \times \alpha)\%$ confidence interval for B(i, l), for i = 1, 2, ..., IP and l = 1, 2, ..., NTAU.

If **Interval Method** = NONE, BU is not referenced.

The method used for calculating the interval is controlled by the optional parameters **Interval Method** and **Bootstrap Interval Method**. The size of the interval, α is controlled by the optional parameter **Significance Level**.

18: CH(IP, IP, *) - REAL (KIND=nag_wp) array

Output

Note: the last dimension of the array CH must be at least NTAU if **Interval Method** \neq NONE and **Matrix Returned** = COVARIANCE and at least NTAU + 1 if **Interval Method** \neq NONE, IID or BOOTSTRAP XY and **Matrix Returned** = H INVERSE.

On exit: depending on the supplied optional parameters, CH will either not be referenced, hold an estimate of the upper triangular part of the covariance matrix, Σ , or an estimate of the upper triangular parts of nJ_n and $n^{-1}H_n^{-1}$.

If Interval Method = NONE or Matrix Returned = NONE, CH is not referenced.

If Interval Method = BOOTSTRAP XY or IID and Matrix Returned = H INVERSE, CH is not referenced.

Otherwise, for $i, j = 1, 2, \dots, IP, j \ge i$ and $l = 1, 2, \dots, NTAU$:

If **Matrix Returned** = COVARIANCE, CH(i, j, l) holds an estimate of the covariance between B(i, l) and B(j, l).

If **Matrix Returned** = H INVERSE, CH(i, j, 1) holds an estimate of the (i, j)th element of nJ_n and CH(i, j, l+1) holds an estimate of the (i, j)th element of $n^{-1}H_n^{-1}$, for $\tau = TAU(l)$.

The method used for calculating Σ and H_n^{-1} is controlled by the optional parameter **Interval Method**.

19: RES(N, *) - REAL (KIND=nag wp) array

Output

Note: the second dimension of the array RES must be at least NTAU if **Return Residuals** = YES.

On exit: if **Return Residuals** = YES, RES(i, l) holds the (weighted) residuals, r_i , for $\tau = \text{TAU}(l)$, for i = 1, 2, ..., N and l = 1, 2, ..., NTAU.

If WEIGHT = 'W' and **Drop Zero Weights** = YES, the value of RES will be set to zero for observations with zero weights.

If **Return Residuals** = NO, RES is not referenced.

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20: IOPTS(*) - INTEGER array

Communication Array

Note: the dimension of this array is dictated by the requirements of associated functions that must have been previously called. This array **must** be the same array passed as argument IOPTS in the previous call to G02ZKF.

On entry: optional parameter array, as initialized by a call to G02ZKF.

21: OPTS(*) - REAL (KIND=nag wp) array

Communication Array

Note: the dimension of this array is dictated by the requirements of associated functions that must have been previously called. This array **must** be the same array passed as argument OPTS in the previous call to G02ZKF.

On entry: optional parameter array, as initialized by a call to G02ZKF.

22: STATE(*) – INTEGER array

Communication Array

Note: the actual argument supplied **must** be the array STATE supplied to the initialization routines G05KFF or G05KGF.

The actual argument supplied **must** be the array STATE supplied to the initialization routines G05KFF or G05KGF.

If **Interval Method** = BOOTSTRAP XY, STATE contains information about the selected random number generator. Otherwise STATE is not referenced.

23: INFO(NTAU) - INTEGER array

Output

On exit: INFO(i) holds additional information concerning the model fitting and confidence limit calculations when $\tau = \text{TAU}(i)$.

Code Warning

- 0 Model fitted and confidence limits (if requested) calculated successfully
- The routine did not converge. The returned values are based on the estimate at the last iteration. Try increasing **Iteration Limit** whilst calculating the parameter estimates or relaxing the definition of convergence by increasing **Tolerance**.
- A singular matrix was encountered during the optimization. The model was not fitted for this value of τ .
- Some truncation occurred whilst calculating the confidence limits for this value of τ . See Section 11 for details. The returned upper and lower limits may be narrower than specified.
- The routine did not converge whilst calculating the confidence limits. The returned limits are based on the estimate at the last iteration. Try increasing **Iteration Limit**.
- 16 Confidence limits for this value of τ could not be calculated. The returned upper and lower limits are set to a large positive and large negative value respectively as defined by the optional parameter **Big**.

It is possible for multiple warnings to be applicable to a single model. In these cases the value returned in INFO is the sum of the corresponding individual nonzero warning codes.

24: IFAIL - INTEGER

Input/Output

On entry: IFAIL must be set to 0, -1 or 1. If you are unfamiliar with this argument you should refer to Section 3.4 in How to Use the NAG Library and its Documentation for details.

For environments where it might be inappropriate to halt program execution when an error is detected, the value -1 or 1 is recommended. If the output of error messages is undesirable, then the value 1 is recommended. Otherwise, if you are not familiar with this argument, the recommended value is 0. When the value -1 or 1 is used it is essential to test the value of IFAIL on exit.

On exit: IFAIL = 0 unless the routine detects an error or a warning has been flagged (see Section 6).

6 Error Indicators and Warnings

If on entry IFAIL = 0 or -1, explanatory error messages are output on the current error message unit (as defined by X04AAF).

Errors or warnings detected by the routine:

```
IFAIL = 11
        On entry, SORDER = \langle value \rangle.
        Constraint: SORDER = 1 or 2.
IFAIL = 21
        On entry, INTCPT = \langle value \rangle was an illegal value.
IFAIL = 31
        On entry, WEIGHT had an illegal value.
IFAIL = 41
        On entry, N = \langle value \rangle.
        Constraint: N \ge 2.
IFAIL = 51
        On entry, M = \langle value \rangle.
        Constraint: M \ge 0.
IFAIL = 71
        On entry, LDDAT = \langle value \rangle and N = \langle value \rangle.
        Constraint: LDDAT \geq N.
IFAIL = 72
        On entry, LDDAT = \langle value \rangle and M = \langle value \rangle.
        Constraint: LDDAT > M.
IFAIL = 81
        On entry, ISX(\langle value \rangle) = \langle value \rangle.
        Constraint: ISX(i) = 0 or 1 for all i.
IFAIL = 91
        On entry, IP = \langle value \rangle and N = \langle value \rangle.
        Constraint: 1 \le IP < N.
IFAIL = 92
        On entry, IP is not consistent with ISX or INTCPT: IP = \langle value \rangle, expected value = \langle value \rangle.
IFAIL = 111
        On entry, WT(\langle value \rangle) = \langle value \rangle.
        Constraint: WT(i) \ge 0.0 for all i.
IFAIL = 112
        On entry, effective number of observations = \langle value \rangle.
        Constraint: effective number of observations \geq \langle value \rangle.
```

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IFAIL = 121

On entry, NTAU = $\langle value \rangle$. Constraint: NTAU ≥ 1 .

IFAIL = 131

On entry, $TAU(\langle value \rangle) = \langle value \rangle$ is invalid.

IFAIL = 201

On entry, either the option arrays have not been initialized or they have been corrupted.

IFAIL = 221

On entry, STATE vector has been corrupted or not initialized.

IFAIL = 231

A potential problem occurred whilst fitting the model(s). Additional information has been returned in INFO.

IFAIL = -99

An unexpected error has been triggered by this routine. Please contact NAG.

See Section 3.9 in How to Use the NAG Library and its Documentation for further information.

IFAIL = -399

Your licence key may have expired or may not have been installed correctly.

See Section 3.8 in How to Use the NAG Library and its Documentation for further information.

IFAIL = -999

Dynamic memory allocation failed.

See Section 3.7 in How to Use the NAG Library and its Documentation for further information.

7 Accuracy

Not applicable.

8 Parallelism and Performance

G02QGF is threaded by NAG for parallel execution in multithreaded implementations of the NAG Library.

G02QGF makes calls to BLAS and/or LAPACK routines, which may be threaded within the vendor library used by this implementation. Consult the documentation for the vendor library for further information.

Please consult the X06 Chapter Introduction for information on how to control and interrogate the OpenMP environment used within this routine. Please also consult the Users' Note for your implementation for any additional implementation-specific information.

9 Further Comments

G02QGF allocates internally approximately the following elements of real storage: $13n + np + 3p^2 + 6p + 3(p+1) \times \text{NTAU}$. If **Interval Method** = BOOTSTRAP XY then a further np elements are required, and this increases by $p \times \text{NTAU} \times \text{Bootstrap Iterations}$ if **Bootstrap Interval Method** = QUANTILE. Where possible, any user-supplied output arrays are used as workspace and so the amount actually allocated may be less. If SORDER = 2, WEIGHT = 'U',

INTCPT = 'N' and IP = M an internal copy of the input data is avoided and the amount of locally allocated memory is reduced by np.

10 Example

A quantile regression model is fitted to Engels 1857 study of household expenditure on food. The model regresses the dependent variable, household food expenditure, against two explanatory variables, a column of ones and household income. The model is fit for five different values of τ and the covariance matrix is estimated assuming Normal IID errors. Both the covariance matrix and the residuals are returned.

10.1 Program Text

```
Program q02qqfe
1
     G02QGF Example Program Text
!
     Mark 26 Release. NAG Copyright 2016.
      .. Use Statements ..
     Use nag_library, Only: g02qgf, g02zkf, g02zlf, g05kff, naq_wp
1
      .. Implicit None Statement ..
     Implicit None
1
      .. Parameters ..
                                       :: lseed = 1, nin = 5, nout = 6
     Integer, Parameter
!
      .. Local Scalars ..
     Real (Kind=nag_wp)
                                       :: df, rvalue
                                        :: genid, i, ifail, ip, ivalue, j, l,
     Integer
                                           ldbl, lddat, ldres, liopts, lopts,
                                           lstate, lwt, m, n, ntau, optype,
                                           sorder, subid, tdch
                                       :: c1, weight
     Character (1)
      Character (30)
                                       :: cvalue, semeth
     Character (100)
                                       :: optstr
!
      .. Local Arrays ..
     Real (Kind=nag_wp), Allocatable :: b(:,:), b1(:,:), bu(:,:), ch(:,:,:), &
                                           dat(:,:), opts(:), res(:,:), tau(:), &
                                           wt(:), y(:)
      Integer, Allocatable
                                        :: info(:), iopts(:), isx(:), state(:)
     Integer
                                        :: seed(lseed)
!
      .. Intrinsic Procedures ..
     Intrinsic
                                        :: count, len_trim, min
!
      .. Executable Statements ..
     Write (nout,*) 'G02QGF Example Program Results'
     Write (nout,*)
     Flush (nout)
     Skip heading in data file
     Read (nin,*)
!
     Read in the problem size
     Read (nin,*) sorder, c1, weight, n, m, ntau
!
     Read in the data
      If (weight=='W' .Or. weight=='w') Then
        lwt = n
     Else
        lwt = 0
     End If
     Allocate (wt(lwt), isx(m), y(n), tau(ntau))
     If (sorder==1) Then
!
       DAT(N,M)
        lddat = n
        Allocate (dat(lddat,m))
        If (lwt==0) Then
         Read (nin,*)(dat(i,1:m),y(i),i=1,n)
```

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```
Read (nin,*)(dat(i,1:m),y(i),wt(i),i=1,n)
        End If
     Else
        DAT(M,N)
        lddat = m
        Allocate (dat(lddat,n))
        If (lwt==0) Then
          Read (nin,*)(dat(1:m,i),y(i),i=1,n)
        Else
         Read (nin,*)(dat(1:m,i),y(i),wt(i),i=1,n)
        End If
     End If
     Read in variable inclusion flags
!
     Read (nin,*) isx(1:m)
     Calculate IP
      ip = count(isx(1:m)==1)
      If (c1=='Y' .Or. c1=='y') Then
        ip = ip + 1
     End If
     Read in the quantiles required
     Read (nin,*) tau(1:ntau)
      liopts = 100
      lopts = 100
     Allocate (iopts(liopts), opts(lopts))
      Initialize the optional argument array
      ifail = 0
      Call g02zkf('INITIALIZE = G02QGF',iopts,liopts,opts,lopts,ifail)
c_lp: Do
        Read in any optional arguments. Reads in to the end of
        the input data, or until a blank line is reached
        ifail = 1
        Read (nin,99994, Iostat=ifail) optstr
        If (ifail/=0) Then
          Exit c_lp
        Else If (len\_trim(optstr)==0) Then
         Exit c_lp
        End If
!
        Set the supplied option
        ifail = 0
        Call g02zkf(optstr,iopts,liopts,opts,lopts,ifail)
     End Do c_lp
     Assume that no intervals or output matrices are required
      unless the optional argument state differently
      ldbl = 0
      tdch = 0
      ldres = 0
     lstate = 0
      Query the optional arguments to see what output is required
      ifail = 0
      Call gO2zlf('INTERVAL METHOD', ivalue, rvalue, cvalue, optype, iopts, opts,
       ifail)
      semeth = cvalue
      If (semeth/='NONE') Then
        Require the intervals to be output
!
        ldbl = ip
        If (semeth=='BOOTSTRAP XY') Then
1
          Need to find the length of the state array for the random
!
          number generator
          Read in the generator ID and a seed
          Read (nin,*) genid, subid, seed(1)
```

```
!
          Query the length of the state array
          Allocate (state(lstate))
          ifail = 0
          Call g05kff(genid, subid, seed, lseed, state, lstate, ifail)
          Deallocate STATE so that it can reallocated later
!
          Deallocate (state)
        End If
        ifail = 0
        Call g02zlf('MATRIX RETURNED',ivalue,rvalue,cvalue,optype,iopts,opts, &
          ifail)
        If (cvalue=='COVARIANCE') Then
          tdch = ntau
        Else If (cvalue=='H INVERSE') Then
          If (semeth=='BOOTSTRAP XY' .Or. semeth=='IID') Then
1
            NB: If we are using bootstrap or IID errors then any request for
1
            H INVERSE is ignored
            tdch = 0
          Else
            tdch = ntau + 1
         End If
        End If
        ifail = 0
        Call gO2zlf('RETURN RESIDUALS', ivalue, rvalue, cvalue, optype, iopts, opts, &
          ifail)
        If (cvalue=='YES') Then
          ldres = n
        End If
     End If
!
     Allocate memory for output arrays
      Allocate (b(ip,ntau),info(ntau),bl(ldbl,ntau),bu(ldbl,ntau),
        ch(ldbl,ldbl,tdch),state(lstate),res(ldres,ntau))
      If (lstate>0) Then
!
        Doing bootstrap, so initialize the RNG
        ifail = 0
        Call g05kff(genid, subid, seed, lseed, state, lstate, ifail)
     End If
!
     Call the model fitting routine
      ifail = -1
      Call g02qgf(sorder,c1,weight,n,m,dat,lddat,isx,ip,y,wt,ntau,tau,df,b,bl, &
        bu,ch,res,iopts,opts,state,info,ifail)
      If (ifail/=0) Then
        If (ifail==231) Then
         Write (nout,*) 'Additional error information (INFO): ', info(1:ntau)
        Else
          Go To 100
        End If
     End If
     Display the parameter estimates
!
      Do l = 1, ntau
        Write (nout, 99999) 'Quantile: ', tau(1)
        Write (nout,*)
        If (ldbl>0) Then
          Write (nout,*) '
                                  Lower
                                           Parameter
                                                       Upper'
          Write (nout,*)'
                                  Limit
                                           Estimate
                                                       Limit'
        Else
          Write (nout,*)'
                                Parameter'
         Write (nout,*) '
                                Estimate'
        End If
        Do j = 1, ip
          If (1db1>0) Then
            Write (nout,99998) j, bl(j,1), b(j,1), bu(j,1)
            Write (nout, 99998) j, b(j,1)
```

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```
End If
        End Do
        Write (nout,*)
        If (tdch==ntau) Then
          Write (nout,*) 'Covariance matrix'
          Do i = 1, ip
            Write (nout, 99997) ch(1:i,i,l)
          End Do
          Write (nout,*)
        Else If (tdch==ntau+1) Then
          Write (nout,*) 'J'
          Do i = 1, ip
            Write (nout, 99997) ch(1:i,i,1)
          End Do
          Write (nout,*)
          Write (nout,*) 'H inverse'
          Do i = 1, ip
            Write (nout, 99997) ch(1:i,i,l+1)
          End Do
          Write (nout,*)
        End If
        Write (nout,*)
     End Do
      If (ldres>0) Then
        Write (nout,*) 'First 10 Residuals'
        Write (nout,*) '
                                                        Quantile'
        Write (nout, 99995) 'Obs.', tau(1:ntau)
        Do i = 1, min(n,10)
          Write (nout, 99996) i, res(i,1:ntau)
        End Do
     Else
        Write (nout,*) 'Residuals not returned'
     End If
     Write (nout,*)
     Continue
99999 Format (1X,A,F6.3)
99998 Format (1X,I3,3(3X,F7.3))
99997 Format (1X,10(E10.3,3X))
99996 Format (2X,I3,10(1X,F10.5))
99995 Format (1X,A,10(3X,F6.3,2X))
99994 Format (A100)
    End Program g02qgfe
```

10.2 Program Data

```
GO2QGF Example Program Data
1 'Ŷ' 'U' 235 1 5
                          :: SORDER, C1, WEIGHT, N, M, NTAU
                        800.7990 572.0807 643.3571 459.8177
1245.6964 907.3969 2551.6615 863.9199
420.1577
          255.8394
 541.4117
          310.9587
                       1245.6964
                                                          863.9199
                                             1795.3226 831.4407
901.1575
         485.6800
                       1201.0002 811.5776
 639.0802 402.9974
                        634.4002
                                  427.7975
                                              1165.7734 534.7610
                                                          392.0502
750.8756
          495.5608
                                  649.9985
                                                815.6212
                        956.2315
                                               1264.2066
 945.7989
          633.7978
                       1148.6010
                                  860.6002
                                                          934.9752
829.3979 630.7566
                       1768.8236 1143.4211
                                              1095.4056 813.3081
                       2822.5330 2032.6792
979.1648 700.4409
                                                447.4479 263.7100
                        922.3548 590.6183 1178.9742
1309.8789 830.9586
                                                          769.0838
1492.3987
          815.3602
                       2293.1920 1570.3911
                                                975.8023
                                                          630.5863
                                              1017.8522
 502.8390
          338.0014
                        627.4726
                                  483.4800
                                                          645.9874
616.7168 412.3613
                        889.9809
                                  600.4804
                                                423.8798
                                                          319.5584
790.9225
          520.0006
                       1162.2000
                                  696.2021
                                                558.7767
                                                          348.4518
555.8786
                                                          614.5068
          452.4015
                       1197.0794
                                  774.7962
                                                943.2487
          512.7201
                        530.7972
                                  390.5984
                                               1348.3002
                                                          662.0096
713.4412
                                               2340.6174 1504.3708
838.7561
                                  612.5619
          658.8395
                       1142.1526
                                                587.1792
 535.0766
          392.5995
                       1088.0039
                                  708.7622
                                                          406.2180
 596.4408
                        484.6612 296.9192
                                               1540.9741
         443.5586
                                                          692.1689
 924.5619
          640.1164
                       1536.0201 1071.4627
                                               1115.8481
                                                          588.1371
487.7583 333.8394
                                              1044.6843 511.2609
                        678.8974 496.5976
```

```
692.6397
           466.9583
                          671.8802
                                    503.3974
                                                  1389.7929
                                                              700.5600
 997.8770
           543.3969
                          690.4683
                                     357.6411
                                                  2497.7860 1301.1451
 506.9995
           317.7198
                          860.6948
                                    430.3376
                                                  1585.3809
                                                              879.0660
 654.1587
           424.3209
                          873.3095
                                     624.6990
                                                   1862.0438
                                                              912.8851
 933.9193
                          894.4598
                                                  2008.8546 1509.7812
           518.9617
                                    582.5413
                                                              484.0605
 433.6813
           338.0014
                         1148.6470
                                     580.2215
                                                    697.3099
 587.5962
           419.6412
                          926.8762
                                     543.8807
                                                    571.2517
                                                              399.6703
                          839.0414
                                                    598.3465
 896.4746
           476.3200
                                     588.6372
                                                              444.1001
                          829.4974
                                                    461.0977
 454.4782
           386.3602
                                     627.9999
                                                              248.8101
 584.9989
           423.2783
                         1264.0043
                                    712.1012
                                                    977.1107
                                                              527.8014
                                                    883.9849
 800.7990
           503.3572
                         1937.9771
                                    968.3949
                                                              500.6313
 502.4369
                                     482.5816
                                                    718.3594
           354.6389
                          698.8317
                                                              436.8107
 713.5197
           497.3182
                          920.4199
                                     593.1694
                                                    543.8971
                                                              374.7990
                         1897.5711 1033.5658
                                                  1587.3480
 906.0006
           588.5195
                                                              726.3921
 880.5969
                                                  4957.8130 1827.2000
           654.5971
                          891.6824
                                     693.6795
 796.8289
           550.7274
                          889.6784
                                     693.6795
                                                    969.6838
                                                              523.4911
                         1221.4818
                                                    419.9980
 854.8791
           528.3770
                                     761.2791
                                                              334.9998
           640.4813
                                     361.3981
                                                    561.9990
1167.3716
                          544.5991
                                                              473.2009
 523.8000
           401.3204
                         1031.4491
                                     628.4522
                                                    689.5988
                                                              581.2029
 670.7792
           435.9990
                         1462.9497
                                     771.4486
                                                  1398.5203
                                                              929.7540
 377.0584
           276.5606
                          830.4353
                                    757.1187
                                                   820.8168
                                                              591.1974
 851.5430
           588.3488
                          975.0415
                                     821.5970
                                                    875.1716
                                                              637.5483
                         1337.9983 1022.3202
                                                  1392.4499
                                                              674.9509
1121.0937
           664.1978
 625.5179
           444.8602
                          867.6427
                                     679.4407
                                                  1256.3174
                                                              776.7589
                          725.7459
                                    538.7491
                                                              959.5170
 805.5377
           462.8995
                                                  1362.8590
                          989.0056
                                                  1999.2552 1250.9643
 558.5812
           377.7792
                                     679.9981
 884.4005
           553.1504
                         1525.0005
                                     977.0033
                                                  1209.4730
                                                              737.8201
                          672.1960
                                                  1125.0356
1257.4989
           810.8962
                                     561.2015
                                                              810.6772
2051.1789 1067.9541
                          923.3977
                                     728.3997
                                                  1827.4010
                                                              983.0009
1466.3330 1049.8788
                          472.3215
                                     372.3186
                                                  1014.1540
                                                              708.8968
                          590.7601
 730.0989
           522.7012
                                     361.5210
                                                    880.3944
                                                              633.1200
2432.3910 1424.8047
                          831.7983
                                     620.8006
                                                    873.7375
                                                              631.7982
 940.9218
           517.9196
                         1139.4945
                                    819.9964
                                                    951.4432
                                                              608.6419
1177.8547
                          507.5169
                                     360.8780
                                                    473.0022
                                                              300.9999
           830.9586
1222.5939
           925.5795
                          576.1972
                                     395.7608
                                                    601.0030
                                                              377.9984
1519.5811 1162.0024
                          696.5991
                                     442.0001
                                                    713.9979
                                                              397.0015
                          650.8180
                                                    829.2984
 687.6638
           383.4580
                                     404.0384
                                                              588.5195
                          949.5802
                                     670.7993
                                                    959.7953
 953.1192
           621.1173
                                                              681.7616
 953.1192
           621.1173
                          497.1193
                                     297.5702
                                                  1212.9613
                                                              807.3603
 953.1192
           621.1173
                          570.1674
                                     353.4882
                                                    958.8743
                                                              696.8011
 939.0418
           548.6002
                          724.7306
                                     383.9376
                                                  1129.4431
                                                              811.1962
1283.4025
           745.2353
                          408.3399
                                    284.8008
                                                  1943.0419 1305.7201
                                                    539.6388
1511.5789
           837.8005
                          638.6713
                                    431.1000
                                                              442.0001
1342.5821
           795.3402
                         1225.7890
                                    801.3518
                                                    463.5990
                                                              353.6013
                          715.3701
 511.7980
           418.5976
                                     448.4513
                                                    562.6400
                                                              468.0008
 689.7988
                                                   736.7584
           508.7974
                          800.4708
                                     577.9111
                                                              526.7573
1532.3074
                          975.5974
                                                   1415.4461
                                                              890.2390
           883.2780
                                    570.5210
1056.0808
           742.5276
                         1613.7565
                                     865.3205
                                                   2208.7897 1318.8033
 387.3195
                                     444.5578
           242.3202
                          608.5019
                                                    636.0009
                                                              331.0005
                                     680.4198
 387.3195
           242.3202
                          958.6634
                                                    759.4010
                                                              416.4015
 410.9987
           266.0010
                          835.9426
                                     576.2779
                                                   1078.8382
                                                              596.8406
           408.4992
                         1024.8177
                                     708.4787
 499.7510
                                                    748.6413
                                                              429.0399
 832.7554
           614.7588
                         1006.4353
                                     734.2356
                                                    987.6417
                                                              619.6408
                          726.0000
                                                    788.0961
 614.9986
           385.3184
                                    433.0010
                                                              400.7990
 887.4658
           515.6200
                          494.4174
                                     327.4188
                                                   1020.0225
                                                              775.0209
                          776.5958
                                                   1230.9235
1595.1611 1138.1620
                                    485.5198
                                                              772.7611
                          415.4407
1807.9520
           993.9630
                                     305.4390
                                                    440.5174
                                                              306.5191
                          581.3599
                                    468.0008
 541.2006
           299.1993
                                                    743.0772
                                                              522.6019
1057.6767
           750.3202
                          :: End of X,Y (in three set of columns)
                          :: ISX
0.10 0.25 0.50 0.75 0.90 :: TAU
Return Residuals = Yes
Matrix Returned = Covariance
Interval Method = IID
```

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10.3 Program Results

G02QGF Example Program Results

Quantile: 0.100

Lower Parameter Upper
Limit Estimate Limit
1 74.946 110.142 145.337
2 0.370 0.402 0.433

Covariance matrix

0.319E+03

-0.254E+00 0.259E-03

Quantile: 0.250

Lower Parameter Upper
Limit Estimate Limit
1 64.232 95.483 126.735
2 0.446 0.474 0.502

Covariance matrix

0.252E+03

-0.200E+00 0.204E-03

Quantile: 0.500

Lower Parameter Upper
Limit Estimate Limit
1 55.399 81.482 107.566
2 0.537 0.560 0.584

Covariance matrix

0.175E+03

-0.140E+00 0.142E-03

Quantile: 0.750

Lower Parameter Upper
Limit Estimate Limit
1 41.372 62.396 83.421
2 0.625 0.644 0.663

Covariance matrix

0.114E+03

-0.907E-01 0.923E-04

Quantile: 0.900

Lower Parameter Upper
Limit Estimate Limit
1 26.829 67.351 107.873
2 0.650 0.686 0.723

Covariance matrix

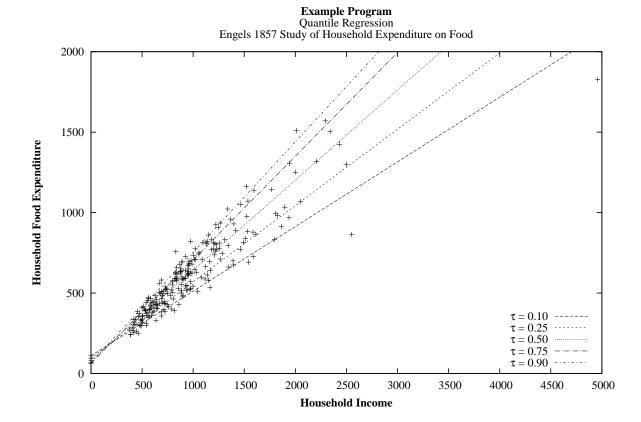
0.423E+03

-0.337E+00 0.343E-03

First 10 Residuals

Obs. 0.100 0.250 0.500 0.750 0.900 1 -23.10718 -38.84219 -61.00711 -77.14462 -99.86551 2 140.20549 96.93582 42.00636 -6.04177 -44.85812 3 91.19725 59.31654 17.93924 -16.90993 -49.06884 4 -16.70358 -41.20981 -73.81193 -100.11463 -127.96277

```
5 296.77717 221.32470 128.09970 42.75414 -14.87476
6 -271.39185 -441.31464 -646.95350 -841.78309 -954.63488
7 13.48419 -37.04518 -100.61322 -157.07478 -200.13481
8 218.91527 146.69601 57.31834 -24.28017 -80.01908
9 0.00000 -115.21109 -255.74639 -387.16920 -468.03911
10 36.09526 4.52393 -36.48522 -70.97584 -102.95390
```



11 Algorithmic Details

By the addition of slack variables the minimization (1) can be reformulated into the linear programming problem

$$\underset{(u,v,\beta)\in\mathbb{R}_{+}^{n}\times\mathbb{R}_{+}^{n}\times\mathbb{R}_{+}^{p}}{\text{minimize}} \tau e^{\mathsf{T}}u + (1-\tau)e^{\mathsf{T}}v \quad \text{subject to} \quad y = X\beta + u - v$$
(2)

and its associated dual

$$\underset{d}{\operatorname{maximize}} y^{\mathsf{T}} d \quad \text{ subject to } \quad X^{\mathsf{T}} d = 0, d \in [\tau - 1, \tau]^n \tag{3}$$

where e is a vector of n 1s. Setting $a = d + (1 - \tau)e$ gives the equivalent formulation

$$\underset{a}{\operatorname{maximize}} y^{\mathsf{T}} a \quad \text{ subject to } \quad X^{\mathsf{T}} a = (1 - \tau) X^{\mathsf{T}} e, a \in [0, 1]^n. \tag{4}$$

The algorithm introduced by Portnoy and Koenker (1997) and used by G02QGF, uses the primal-dual formulation expressed in equations (2) and (4) along with a logarithmic barrier function to obtain estimates for β . The algorithm is based on the predictor-corrector algorithm of Mehrotra (1992) and further details can be obtained from Portnoy and Koenker (1997) and Koenker (2005). A good description of linear programming, interior point algorithms, barrier functions and Mehrotra's predictor-corrector algorithm can be found in Nocedal and Wright (1999).

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11.1 Interior Point Algorithm

In this section a brief description of the interior point algorithm used to estimate the model parameters is presented. It should be noted that there are some differences in the equations given here – particularly (7) and (9) – compared to those given in Koenker (2005) and Portnoy and Koenker (1997).

11.1.1 Central path

Rather than optimize (4) directly, an additional slack variable s is added and the constraint $a \in [0, 1]^n$ is replaced with $a + s = e, a_i \ge 0, s_i \ge 0$, for i = 1, 2, ..., n.

The positivity constraint on a and s is handled using the logarithmic barrier function

$$B(a,s,\mu) = y^{\mathsf{T}}a + \mu \sum_{i=1}^n (\log a_i + \log s_i).$$

The primal-dual form of the problem is used giving the Lagrangian

$$L(a, s, \beta, u, \mu) = B(a, s, \mu) - \beta^{T} (X^{T} a - (1 - \tau) X^{T} e) - u^{T} (a + s - e)$$

whose central path is described by the following first order conditions

$$X^{\mathsf{T}}a = (1-\tau)X^{\mathsf{T}}e$$

$$a+s = e$$

$$X\beta + u - v = y$$

$$SUe = \mu e$$

$$AVe = \mu e$$

$$(5)$$

where A denotes the diagonal matrix with diagonal elements given by a, similarly with S, U and V. By enforcing the inequalities on s and a strictly, i.e., $a_i > 0$ and $s_i > 0$ for all i we ensure that A and S are positive definite diagonal matrices and hence A^{-1} and S^{-1} exist.

Rather than applying Newton's method to the system of equations given in (5) to obtain the step directions δ_{β} , δ_{a} , δ_{s} , δ_{u} and δ_{v} , Mehrotra substituted the steps directly into (5) giving the augmented system of equations

$$X^{\mathsf{T}}(a+\delta_a) = (1-\tau)X^{\mathsf{T}}e$$

$$(a+\delta_a) + (s+\delta_s) = e$$

$$X(\beta+\delta_\beta) + (u+\delta_u) - (v+\delta_v) = y$$

$$(S+\Delta_s)(U+\Delta_u)e = \mu e$$

$$(A+\Delta_a)(V+\Delta_v)e = \mu e$$

$$(A+\Delta_a)(V+\Delta_v)e = \mu e$$

$$(6)$$

where $\Delta_a, \Delta_s, \Delta_u$ and Δ_v denote the diagonal matrices with diagonal elements given by $\delta_a, \delta_s, \delta_u$ and δ_v respectively.

11.1.2 Affine scaling step

The affine scaling step is constructed by setting $\mu = 0$ in (5) and applying Newton's method to obtain an intermediate set of step directions

$$(X^{\mathsf{T}}WX)\delta_{\beta} = X^{\mathsf{T}}W(y - X\beta) + (\tau - 1)X^{\mathsf{T}}e + X^{\mathsf{T}}a$$

$$\delta_{a} = W(y - X\beta - X\delta_{\beta})$$

$$\delta_{s} = -\delta_{a}$$

$$\delta_{u} = S^{-1}U\delta_{a} - Ue$$

$$\delta_{v} = A^{-1}V\delta_{s} - Ve$$

$$(7)$$

where $W = (S^{-1}U + A^{-1}V)^{-1}$

Initial step sizes for the primal $(\hat{\gamma}_P)$ and dual $(\hat{\gamma}_D)$ parameters are constructed as

$$\hat{\gamma}_{P} = \sigma \quad \min \left\{ \min_{i, \delta_{a_{i}} < 0} \left\{ a_{i} / \delta_{a_{i}} \right\}, \min_{i, \delta_{s_{i}} < 0} \left\{ s_{i} / \delta_{s_{i}} \right\} \right\}
\hat{\gamma}_{D} = \sigma \quad \min \left\{ \min_{i, \delta_{u_{i}} < 0} \left\{ u_{i} / \delta_{u_{i}} \right\}, \min_{i, \delta_{v_{i}} < 0} \left\{ v_{i} / \delta_{v_{i}} \right\} \right\}$$
(8)

where σ is a user-supplied scaling factor. If $\hat{\gamma}_P \times \hat{\gamma}_D \ge 1$ then the nonlinearity adjustment, described in Section 11.1.3, is not made and the model parameters are updated using the current step size and directions.

11.1.3 Nonlinearity Adjustment

In the nonlinearity adjustment step a new estimate of μ is obtained by letting

$$\hat{g}(\hat{\gamma}_P, \hat{\gamma}_D) = (s + \hat{\gamma}_P \delta_s)^{\mathsf{T}} (u + \hat{\gamma}_D \delta_u) + (a + \hat{\gamma}_P \delta_a)^{\mathsf{T}} (v + \hat{\gamma}_D \delta_v)$$

and estimating μ as

$$\mu = \left(\frac{\hat{g}\left(\hat{\gamma}_{P}, \hat{\gamma}_{D}\right)}{\hat{g}\left(0, 0\right)}\right)^{3} \frac{\hat{g}\left(0, 0\right)}{2n}.$$

This estimate, along with the nonlinear terms (Δu , Δs , Δa and Δv) from (6) are calculated using the values of δ_a , δ_s , δ_u and δ_v obtained from the affine scaling step.

Given an updated estimate for μ and the nonlinear terms the system of equations

$$(X^{\mathsf{T}}WX)\delta_{\beta} = X^{\mathsf{T}}W(y - X\beta + \mu(S^{-1} - A^{-1})e + S^{-1}\Delta_{s}\Delta_{u}e - A^{-1}\Delta_{a}\Delta_{v}e) + (\tau - 1)X^{\mathsf{T}}e + X^{\mathsf{T}}a$$

$$\delta_{a} = W(y - X\beta - X\delta_{\beta} + \mu(S^{-1} - A^{-1}))$$

$$\delta_{s} = -\delta_{a}$$

$$\delta_{u} = \mu S^{-1}e + S^{-1}U\delta_{a} - Ue - S^{-1}\Delta_{s}\Delta_{u}e$$

$$\delta_{v} = \mu A^{-1}e + A^{-1}V\delta_{s} - Ve - A^{-1}\Delta_{a}\Delta_{v}e$$
(9)

are solved and updated values for $\delta_{\beta}, \delta_{a}, \delta_{s}, \delta_{u}, \delta_{v}, \hat{\gamma}_{P}$ and $\hat{\gamma}_{D}$ calculated.

11.1.4 Update and convergence

At each iteration the model parameters (β, a, s, u, v) are updated using step directions, $(\delta_{\beta}, \delta_{a}, \delta_{s}, \delta_{u}, \delta_{v})$ and step lengths $(\hat{\gamma}_{P}, \hat{\gamma}_{D})$.

Convergence is assessed using the duality gap, that is, the differences between the objective function in the primal and dual formulations. For any feasible point (u, v, s, a) the duality gap can be calculated from equations (2) and (3) as

$$\tau e^{\mathsf{T}} u + (1 - \tau) e^{\mathsf{T}} v - d^{\mathsf{T}} y = \tau e^{\mathsf{T}} u + (1 - \tau) e^{\mathsf{T}} v - (a - (1 - \tau)e)^{\mathsf{T}} y
= s^{\mathsf{T}} u + a^{\mathsf{T}} v
= e^{\mathsf{T}} u - a^{\mathsf{T}} y + (1 - \tau) e^{\mathsf{T}} X \beta$$

and the optimization terminates if the duality gap is smaller than the tolerance supplied in the optional parameter **Tolerance**.

11.1.5 Additional information

Initial values are required for the parameters a, s, u, v and β . If not supplied by the user, initial values for β are calculated from a least squares regression of y on X. This regression is carried out by first constructing the cross-product matrix X^TX and then using a pivoted QR decomposition as performed by F08BFF (DGEQP3). In addition, if the cross-product matrix is not of full rank, a rank reduction is carried out and, rather than using the full design matrix, X, a matrix formed from the first p-rank columns of XP is used instead, where P is the pivot matrix used during the QR decomposition. Parameter estimates, confidence intervals and the rows and columns of the matrices returned in the argument CH (if any) are set to zero for variables dropped during the rank-reduction. The rank reduction step is performed irrespective of whether initial values are supplied by the user.

Once initial values have been obtained for β , the initial values for u and v are calculated from the residuals. If $|r_i| < \epsilon_u$ then a value of $\pm \epsilon_u$ is used instead, where ϵ_u is supplied in the optional parameter **Epsilon**. The initial values for the a and s are always set to $1 - \tau$ and τ respectively.

The solution for δ_{β} in both (7) and (9) is obtained using a Bunch–Kaufman decomposition, as implemented in F07MDF (DSYTRF).

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11.2 Calculation of Covariance Matrix

G02QGF supplies four methods to calculate the covariance matrices associated with the parameter estimates for β . This section gives some additional detail on three of the algorithms, the fourth, (which uses bootstrapping), is described in Section 3.

(i) Independent, identically distributed (IID) errors

When assuming IID errors, the covariance matrices depend on the sparsity, $s(\tau)$, which G02QGF estimates as follows:

- (a) Let r_i denote the residuals from the original quantile regression, that is $r_i = y_i x_i^T \hat{\beta}$.
- (b) Drop any residual where $|r_i|$ is less than ϵ_u , supplied in the optional parameter **Epsilon**.
- (c) Sort and relabel the remaining residuals in ascending order, by absolute value, so that $\epsilon_u < |r_1| < |r_2| < \dots$
- (d) Select the first l values where $l = h_n n$, for some bandwidth h_n .
- (e) Sort and relabel these l residuals again, so that $r_1 < r_2 < \ldots < r_l$ and regress them against a design matrix with two columns (p=2) and rows given by $x_i = \{1, i/(n-p)\}$ using quantile regression with $\tau = 0.5$.
- (f) Use the resulting estimate of the slope as an estimate of the sparsity.
- (ii) Powell Sandwich

When using the Powell Sandwich to estimate the matrix H_n , the quantity

$$c_n = \min(\sigma_r, (q_{r3} - q_{r1})/1.34) \times (\Phi^{-1}(\tau + h_n) - \Phi^{-1}(\tau - h_n))$$

is calculated. Dependent on the value of τ and the method used to calculate the bandwidth (h_n) , it is possible for the quantities $\tau \pm h_n$ to be too large or small, compared to **machine precision** (ϵ) . More specifically, when $\tau - h_n \leq \sqrt{\epsilon}$, or $\tau + h_n \geq 1 - \sqrt{\epsilon}$, a warning flag is raised in INFO, the value is truncated to $\sqrt{\epsilon}$ or $1 - \sqrt{\epsilon}$ respectively and the covariance matrix calculated as usual.

(iii) Hendricks-Koenker Sandwich

The Hendricks-Koenker Sandwich requires the calculation of the quantity $d_i = x_i^{\rm T} \Big(\hat{\beta}(\tau + h_n) - \hat{\beta}(\tau - h_n) \Big)$. As with the Powell Sandwich, in cases where $\tau - h_n \leq \sqrt{\epsilon}$, or $\tau + h_n \geq 1 - \sqrt{\epsilon}$, a warning flag is raised in INFO, the value truncated to $\sqrt{\epsilon}$ or $1 - \sqrt{\epsilon}$ respectively and the covariance matrix calculated as usual.

In addition, it is required that $d_i > 0$, in this method. Hence, instead of using $2h_n/d_i$ in the calculation of H_n , $\max(2h_n/(d_i + \epsilon_u), 0)$ is used instead, where ϵ_u is supplied in the optional parameter **Epsilon**.

12 Optional Parameters

Several optional parameters in G02QGF control aspects of the optimization algorithm, methodology used, logic or output. Their values are contained in the arrays IOPTS and OPTS; these must be initialized before calling G02QGF by first calling G02ZKF with OPTSTR set to **Initialize** = G02QGF.

Each optional parameter has an associated default value; to set any of them to a non-default value, use G02ZKF. The current value of an optional parameter can be queried using G02ZLF.

The remainder of this section can be skipped if you wish to use the default values for all optional parameters.

The following is a list of the optional parameters available. A full description of each optional parameter is provided in Section 12.1.

Band Width Alpha Band Width Method Big

Bootstrap Interval Method

Bootstrap Iterations

Bootstrap Monitoring

Calculate Initial Values

Defaults

Drop Zero Weights

Epsilon

Interval Method

Iteration Limit

Matrix Returned

Monitoring

QR Tolerance

Return Residuals

Sigma

Significance Level

Tolerance

Unit Number

12.1 Description of the Optional Parameters

For each option, we give a summary line, a description of the optional parameter and details of constraints.

The summary line contains:

the keywords, where the minimum abbreviation of each keyword is underlined (if no characters of an optional qualifier are underlined, the qualifier may be omitted);

a parameter value, where the letters a, i and r denote options that take character, integer and real values respectively;

the default value, where the symbol ϵ is a generic notation for *machine precision* (see X02AJF). Keywords and character values are case and white space insensitive.

Band Width Alpha r Default = 1.0

A multiplier used to construct the parameter α_b used when calculating the Sheather-Hall bandwidth (see Section 3), with $\alpha_b = (1 - \alpha) \times \textbf{Band Width Alpha}$. Here, α is the **Significance Level**.

Constraint: Band Width Alpha > 0.0.

Band Width Method a Default = 'SHEATHER HALL'

The method used to calculate the bandwidth used in the calculation of the asymptotic covariance matrix Σ and H^{-1} if **Interval Method** = HKS, KERNEL or IID (see Section 3).

Constraint: Band Width Method = SHEATHER HALL or BOFINGER.

 $\mathbf{Big} \qquad \qquad r \qquad \qquad \mathsf{Default} = 10.0^{20}$

This parameter should be set to something larger than the biggest value supplied in DAT and Y.

Constraint: $\mathbf{Big} > 0.0$.

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Bootstrap Interval Method

a

Default = QUANTILE

If **Interval Method** = BOOTSTRAP XY, **Bootstrap Interval Method** controls how the confidence intervals are calculated from the bootstrap estimates.

Bootstrap Interval Method = T

t intervals are calculated. That is, the covariance matrix, $\Sigma = \left\{\sigma_{ij}: i, j = 1, 2, \ldots, p\right\}$ is calculated from the bootstrap estimates and the limits calculated as $\beta_i \pm t_{(n-p,(1+\alpha)/2)}\sigma_{ii}$ where $t_{(n-p,(1+\alpha)/2)}$ is the $(1+\alpha)/2$ percentage point from a Student's t distribution on n-p degrees of freedom, t is the effective number of observations and t is given by the optional parameter **Significance Level**.

Bootstrap Interval Method = QUANTILE

Quantile intervals are calculated. That is, the upper and lower limits are taken as the $(1 + \alpha)/2$ and $(1 - \alpha)/2$ quantiles of the bootstrap estimates, as calculated using G01AMF.

Constraint: Bootstrap Interval Method = \underline{T} or QUANTILE.

Bootstrap Iterations

i

Default = 100

The number of bootstrap samples used to calculate the confidence limits and covariance matrix (if requested) when **Interval Method** = BOOTSTRAP XY.

Constraint: **Bootstrap Iterations** > 1.

Bootstrap Monitoring

a

Default = NO

If **Bootstrap Monitoring** = YES and **Interval Method** = BOOTSTRAP XY, then the parameter estimates for each of the bootstrap samples are displayed. This information is sent to the unit number specified by **Unit Number**.

Constraint: **Bootstrap Monitoring** = \underline{YES} or \underline{NO} .

Calculate **Initial Values**

a

Default = YES

If **Calculate Initial Values** = YES then the initial values for the regression parameters, β , are calculated from the data. Otherwise they must be supplied in B.

Constraint: Calculate Initial Values = \underline{YES} or \underline{NO} .

Defaults

This special keyword is used to reset all optional parameters to their default values.

Drop Zero Weights

a

Default = YES

If a weighted regression is being performed and **Drop Zero Weights** = YES then observations with zero weight are dropped from the analysis. Otherwise such observations are included.

Constraint: Drop Zero Weights = YES or NO.

Epsilon

r

Default $=\sqrt{\epsilon}$

 ϵ_u , the tolerance used when calculating the covariance matrix and the initial values for u and v. For additional details see Section 11.2 and Section 11.1.5 respectively.

Constraint: **Epsilon** ≥ 0.0 .

Interval Method

a

Default = IID

The value of **Interval Method** controls whether confidence limits are returned in BL and BU and how these limits are calculated. This parameter also controls how the matrices returned in CH are calculated.

Interval Method = NONE

No limits are calculated and BL, BU and CH are not referenced.

Interval Method = KERNEL

The Powell Sandwich method with a Gaussian kernel is used.

Interval Method = HKS

The Hendricks-Koenker Sandwich is used.

Interval Method = IID

The errors are assumed to be identical, and independently distributed.

Interval Method = BOOTSTRAP XY

A bootstrap method is used, where sampling is done on the pair (y_i, x_i) . The number of bootstrap samples is controlled by the parameter **Bootstrap Iterations** and the type of interval constructed from the bootstrap samples is controlled by **Bootstrap Interval Method**.

Constraint: Interval Method = NONE, KERNEL, HKS, IID or BOOTSTRAP XY.

Iteration Limit i Default = 100

The maximum number of iterations to be performed by the interior point optimization algorithm.

Constraint: Iteration Limit > 0.

<u>Matrix Returned</u> a Default = NONE

The value of **Matrix Returned** controls the type of matrices returned in CH. If **Interval Method** = NONE, this parameter is ignored and CH is not referenced. Otherwise:

Matrix Returned = NONE

No matrices are returned and CH is not referenced.

Matrix Returned = COVARIANCE

The covariance matrices are returned.

Matrix Returned = H INVERSE

If Interval Method = KERNEL or HKS, the matrices J and H^{-1} are returned. Otherwise no matrices are returned and CH is not referenced.

The matrices returned are calculated as described in Section 3, with the algorithm used specified by **Interval Method**. In the case of **Interval Method** = BOOTSTRAP XY the covariance matrix is calculated directly from the bootstrap estimates.

Constraint: Matrix Returned = \underline{NONE} , \underline{COV} ARIANCE or \underline{H} \underline{INV} ERSE.

Monitoring a Default = NO

If **Monitoring** = YES then the duality gap is displayed at each iteration of the interior point optimization algorithm. In addition, the final estimates for β are also displayed.

The monitoring information is sent to the unit number specified by Unit Number.

Constraint: **Monitoring** = \underline{YES} or \underline{NO} .

QR Tolerance r Default $= \epsilon^{0.9}$

The tolerance used to calculate the rank, k, of the $p \times p$ cross-product matrix, X^TX . Letting Q be the orthogonal matrix obtained from a QR decomposition of X^TX , then the rank is calculated by comparing Q_{ii} with $Q_{11} \times \mathbf{QR}$ Tolerance.

If the cross-product matrix is rank deficient, then the parameter estimates for the p-k columns with the smallest values of Q_{ii} are set to zero, along with the corresponding entries in BL, BU and CH, if returned. This is equivalent to dropping these variables from the model. Details on the QR decomposition used can be found in F08BFF (DGEQP3).

Constraint: **QR Tolerance** > 0.0.

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Return Residuals a Default = NO

If **Return Residuals** = YES, the residuals are returned in RES. Otherwise RES is not referenced.

Constraint: Return Residuals = \underline{YES} or \underline{NO} .

Sigma r Default = 0.99995

The scaling factor used when calculating the affine scaling step size (see equation (8)).

Constraint: 0.0 < Sigma < 1.0.

Significance Level r Default = 0.95

 α , the size of the confidence interval whose limits are returned in BL and BU.

Constraint: 0.0 < Significance Level < 1.0.

Tolerance r Default $=\sqrt{\epsilon}$

Convergence tolerance. The optimization is deemed to have converged if the duality gap is less than **Tolerance** (see Section 11.1.4).

Constraint: Tolerance > 0.0.

Unit Number i Default taken from X04ABF

The unit number to which any monitoring information is sent.

Constraint: Unit Number > 1.

13 Description of Monitoring Information

See the description of the optional argument Monitoring.

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